Functional impairment, illness burden, and depressive symptoms in older adults: does type of social relationship matter?

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Objective: The nature of interpersonal relationships, whether supportive or critical, may affect the association between health status and mental health outcomes. We examined the potential moderating effects of social support, as a buffer, and family criticism, as an exacerbating factor, on the association between illness burden, functional impairment and depressive symptoms.

Methods: Our sample of 735 older adults, 65 years and older, was recruited from internal and family medicine primary care offices. Trained interviewers administered the Hamilton Rating Scale for Depression, Duke Social Support Inventory, and Family Emotional Involvement and Criticism Scale. Physician-rated assessments of health, including the Karnofsky Performance Status Scale and Cumulative Illness Rating Scale, were also completed.

Results: Linear multivariable hierarchical regression results indicate that social interaction was a significant buffer, weakening the association between illness burden and depressive symptoms, whereas perceived social support buffered the relationship between functional impairment and depressive symptoms. Family criticism and instrumental social support were not significant moderators.

Conclusions: Type of medical dysfunction, whether illness or impairment, may require different therapeutic and supportive approaches. Enhancement of perceived social support, for those who are impaired, and encouragement of social interactions, for those who are ill, may be important intervention targets for treatment of depressive symptoms in older adult primary care patients. Copyright © 2012 John Wiley & Sons, Ltd.

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In the United States, approximately 13% of the adult population is diagnosed with major depressive disorder in their lifetime (Hasin *et al.*, 2005). Of particular concern, older adults may be at greater risk for depression than other age groups because of developmental, role expectation, and health-related changes (Teachman, 2006). Declining physical health is often a primary trigger for depression (Yang, 2006), and epidemiological data suggest that 75% of adults 65 and older have at least one, and almost 50% have two or more, chronic illnesses (National Center for Chronic Disease Prevention and Health Promotion, 2000). The association between chronic medical conditions and depressive symptoms may be tempered, or exacerbated, by the degree and quality of an individual's interpersonal relationships. The Interpersonal Theory of Depression proposes that difficulties in interpersonal relationships, perhaps because of recent changes in a major interpersonal role resulting from illness, are a robust contributing factor to the onset and maintenance of depressive symptoms (Davidson *et al.*, 2004; Weissman, 2010). Having a large, active social network and the receipt of emotional support may provide a buffer against health-related stressors (Penninx *et al.*, 1998; Yang, 2006), whereas loneliness, isolation, or negative social exchanges may exacerbate the effect of poor health on emotional functioning (Cacioppo *et al.*, 2006).

In the current study, we investigated the role of three types of social support as potential buffers of the association between illness, impairment, and depressive symptoms, including instrumental support (i.e., assistance with tasks of daily living), perceived support (i.e., subjective quality of and satisfaction with one's social network), and social interaction (i.e., the actual extent and availability of one's social network) (George *et al.*, 1989; Landerman *et al.*, 1989). Of these, perceived social support is often more predictive of depressive symptoms than the objective forms of support (Antonucci *et al.*, 1997).

Not all interpersonal relationships are positive or helpful in nature. Concerns over close relationships, and negative social exchanges, are associated with elevated depressive symptoms (Krause et al., 1989). Importantly, dysfunctional familial relationships may impact well-being more than conflicts in peripheral relationships (Abbey et al., 1985), as they threaten enduring commitments. In the context of health dysfunction, family members may become critical, disapproving, or rejecting of the health behaviors and decisions of an older adult (Seaburn et al., 2005), which may result in further functional decline, maladaptive health behaviors, increased negative affect, and depression (Shields et al., 1992; Bressi et al., 2007). As such, in the current study, we examined the role of family criticism, or the extent to which an individual perceives the receipt of critical comments from family members (Shields et al., 1994), as a potential exacerbating factor that might strengthen the association between illness or impairment and depressive symptoms.

It is not clear from previous research, however, what type of social support, or negative social interaction, is most salient for illness versus impairment, if any differences exist. We hypothesized that both illness burden and functional impairment would be independently associated with greater levels of depressive symptoms. Furthermore, we examined the potential independent and combined moderating effects of family criticism and social support, hypothesizing that family criticism would exacerbate the association between illness/impairment and depressive symptoms, whereas social support variables would buffer this relationship.

Method

Participants

Patients 65 years and older were recruited from internal medicine and family medicine primary care offices in the Rochester (New York) region, as part of a ongoing, institutional review board-approved study of depression and medical comorbidity; written informed consent was obtained, and no author conflicts of interest exist (Hirsch et al., 2007). Of eligible subjects, 735 (50.1%) completed an interview, which were conducted in their homes or at a university research office by a trained rater. Our sample ranged in age from 65 to 97 years old (mean age = 75.14, SD = 6.89), was 63.4% female (n = 466), and had an average education level of 14.03 years (SD = 2.61). Participants were predominantly White people (92.1%; n = 677) and largely lived alone (34%; n = 251) or with a spouse or significant other (48%; n = 353). Enrolled subjects did not significantly differ from nonenrolled patients in age, gender, or the 15-item Geriatric Depression Scale score. Cognitive status for this sample (mean Mini mental state examination (MMSE) score = 27.61, SD = 2.49) indicates largely intact cognitive functioning (see Table 1 for mean and standard deviation scores).

Measures

Demographic characteristics, including age, sex, education level, ethnicity, living arrangements, and cognitive status, were assessed. Because of the common comorbidity between cognitive status and depressive symptoms in older adults (Potter and Steffens, 2007), we covaried scores on the MMSE (Folstein *et al.*, 1975), a screener for cognitive functioning and impairment (Tombaugh and McIntyre, 1992). The MMSE exhibits adequate reliability and validity in use with older adults (Douglas *et al.*, 2008). Mean score for the current sample was 27.61 (SD = 2.49).

We assessed depressive symptoms utilizing the Hamilton Depression Rating Scale (HDRS), a 24-item structured clinical interview measuring symptom frequency and severity (Hamilton, 1960). The HDRS uses a five-point scale (0 = absence of a symptom), 4 = endorsement of a severe level of a symptom); higher scores indicate more severe symptoms. The HDRS exhibits adequate psychometric properties, including in use with primary care patients (Iannuzzo *et al.*, 2006); in the current study, Cronbach's $\alpha = 0.80$.

	Mean (SD) <i>n</i> (%)	Age	Education	Cognitive status	Illness burden	Depressive symptoms	Social interaction	Perceived social support	Instrumental social support	Family criticism	Functional impairment
Gender (female)	466 (63.4%)	-0.08*	0.18**	-0.07	-0.01	-0.21**	-0.04	0.02	0.06	60.0	-0.14**
Ade	75.14 (6.89)	Ι	-0.11**	-0.26**	0.25**	0.06	-0.10**	-0.04	0.04	-0.10**	0.28**
Education	14.03 (2.61)	I		0.41**	-0.13**	-0.16**	0.25**	0.18**	0.05	-0.09*	-0.24**
Cognitive status	27.61 (2.49)				-0.16**	-0.09*	0.23**	0.18**	0.04	-0.06	-0.26**
Ilness burden	7.55 (3.00)		I			0.33**	-0.17**	-0.17**	-0.07	0.05	0.59**
Depressive symptoms	8.73 (6.05)					I	-0.20**	-0.36**	0.01	0.20**	0.39**
Social interaction	8.96 (1.58)					I	I	0.31**	-0.18**	-0.12**	-0.21**
Perceived social support	19.51 (2.19)					I	I	I	-0.31**	-0.42**	-0.28**
Instrumental social support	14.81 (2.38)					I	I	I	Ι	0.04	-0.04
Family criticism	1.99 (2.95)			I		I	I	I	Ι		0.07
Functional impairment	21.79 (13.20)	Ι				I	I	I	I		

Social support, impairment, illness, and depression

The Duke Social Support Index (DSSI) (Landerman et al., 1989), a 23-item measure employing Likert scales and yes/no answers, was used to assess three components of social support: instrumental support (12 items) (e.g., do friends or family help you out when you are sick?), perceived support (seven items) (e.g., when you are talking to your family and friends do you feel you are being listened to?), and social interaction (four items) (e.g., other than members of your family, how many persons in this area, within one hour of travel can you depend on or feel close to?). In previous research with older adults, the DSSI exhibited adequate internal consistency (Cronbach's $\alpha = 0.77$) (Powers *et al.*, 2004). In the current study, subscale Cronbach's $\alpha = 0.81$ for social interaction, 0.91 for perceived support, and 0.96 for instrumental support.

The Family Emotional Involvement and Criticism Scale (FEICS) (Shields *et al.*, 1992) was used to assess critical comments from family members. Participants responded to the seven-item perceived criticism subscale (e.g., *My family approves of most everything I do; My family complains about what I do for fun*); anchor ratings range from 0 (almost never) to 4 (almost always). The FEICS exhibits sufficient construct and criterion validity, and adequate internal consistency ($\alpha = 0.82$) in use with older adult and primary care patients (Shields *et al.*, 1994). In the current study, Cronbach's $\alpha = 0.72$.

Chronic medical problems were objectively assessed using the Cumulative Illness Rating Scale (CIRS) (Linn *et al.*, 1968), completed by each patients' primary physician, who assessed degree of pathology and impairment present in major organ groups. Physicians used a five-point scale (0 = none, 4 = extremely severe) to assess illness burden in each of six categories: cardiovascular/respiratory, gastrointestinal, genitourinary, musculoskeletal or integument, neuropsychiatric, and general (endocrine metabolic). CIRS scores were derived from laboratory evaluations, physical examinations, and medical history collected from interviews and health records (Hirsch *et al.*, 2007).

Finally, medically induced disability was assessed using the Karnofsky Performance Status Scale (KPSS) (Karnofsky and Barchenal, 1949), a physician-rated scale, which is scored from 0 to 100 (0 = dead, 100 = no evidence of disease). In our analyses, the KPSS is reverse-scored for interpretability; higher scores represent greater impairment. The KPSS exhibits adequate validity in use with geriatric patients and is associated with greater depressive symptoms and suicide risk (Crooks *et al.*, 1991; Conwell *et al.*, 2000).

Criticism Scale.

p < 0.05. **p < 0.01.

Statistical analyses

To assess bivariate associations, Pearson's productmoment correlation coefficients were calculated; no variables met criteria for multicollinearity (Tabachnick and Fidell, 2001). Multivariate, hierarchical linear regressions were used to conduct moderation analyses (Baron and Kenny, 1986), and predictor and moderator variables were centered prior to analyses (Aiken and West, 1991). We covaried age, education, ethnicity, sex, cognitive status, and living arrangements that, along with predictors, were entered on the first step of regression models; interaction terms were entered on the second step. In analyses examining functional impairment, illness burden was statistically controlled, and likewise, functional impairment was controlled in analyses examining illness burden. To assess the individual effects of each potential moderator, we not only tested independent analytic models but also constructed a combined analytic model to ascertain interaction effects within the context of other potential moderators. To create graphic displays of interactions, grouping variables were split one standard deviation higher and lower than the mean, except for perceived social support and family criticism, which were split at the median owing to skew and kurtosis.

Results

Basic associations between variables occurred in expected directions (p < 0.001; Table 1). Illness burden (r=0.33), functional impairment (r=0.39), and family criticism (r=0.20) were significantly positively associated, and social interaction (r=-0.20) and perceived social support (r=-0.36) were significantly negatively associated, with depressive symptoms.

The potential moderating effect of perceived social support on the association between cumulative illness burden and depressive symptoms (standardized $\beta = 0.18$, p = 0.001) was not significant. There was a main effect for perceived social support, which was associated with fewer depressive symptoms (t = -7.83, p < 0.001; standardized $\beta = -0.27$; Un $\beta = -0.77$ (SE = 0.10)).

Higher functional impairment scores were associated with greater levels of depressive symptoms (t = 5.30, p < 0.001; standardized β = 0.21), and perceived social support significantly moderated this relationship ($F_{(1, 679)}$ = 7.71; t = 2.78, p = 0.006). Those with greater perceived support have lower levels of

depressive symptoms related to functional impairment (Table 2 and Figure 1). There were also main effects for perceived support, which was associated with fewer depressive symptoms (t=-8.32, p < 0.001), and for illness burden (t=4.40, p < 0.001), which was associated with greater depressive symptoms. Although our results indicate the presence of an interaction, a graphical review suggests that it is negligible; a main effect seems more predominant, such that although individuals with greater perceived support have less depressive symptoms across levels of functional impairment than those with less perceived support, they remain at increasing risk for depressive symptoms as impairment increases.

Greater illness burden was associated with higher levels of depressive symptoms (t=4.38, p < 0.001; standardized β = 0.18; Un β = 0.37 (SE = 0.09)), and social interaction was a significant moderator ($F_{(1, 674)}$ =6.24, t=-2.50, p < 0.05, Un β =-0.11 (SE = 0.04); R^2 = 0.236, R^2 = 0.007, p < 0.05). Those with more social interaction have lower levels of depressive symptoms associated with illness burden (Table 3 and Figure 2). Main effects existed for functional impairment (t=5.55, p < 0.001; Un β =0.12 (SE = 0.02)), which was associated with greater depressive symptoms, and for social interaction (t=-3.42, p < 0.01; standardized β =-0.13; Un β =-0.47 (SE = 0.14)), which was associated with fewer depressive symptoms.

In an analysis of the potential moderating effect of social interaction on the association between functional impairment and depressive symptoms (t=5.71, p < 0.001; standardized $\beta = 0.26$; Un $\beta = 0.12$ (SE = 0.02)), social interaction was not a significant moderator. However, there was a significant main effect for social interaction, which was associated with fewer depressive symptoms (t=-3.60, p < 0.001; standardized $\beta = -0.13$; Un $\beta = -0.49$ (SE = 0.14)).

Although there were main effects for the associations between cumulative illness burden and depression $(F_{(1, 672)} = 0.00, p = 0.971;$ standardized $\beta = 0.19, p = 0.000$) and functional impairment and depression $(F_{(1, 672)} = 1.67, p = 0.197;$ standardized $\beta = 0.27, p = 0.000$), instrumental social support was not a significant moderator in either analyses. Similarly, in analyses of family criticism, higher illness burden $(F_{(1, 655)} = 0.20, p = 0.656;$ standardized $\beta = 0.18, p = 0.000)$ and functional impairment $(F_{(1, 655)} = 0.53, p = 0.465;$ standardized $\beta = 0.27, p = 0.000$) were associated with more depressive symptoms; however, family criticism was not a significant moderator. There was a main effect for family criticism, which was related to greater

Table 2 Multivariable hierarchical linear regressions of social relationship variables as moderators of association between functional impairment and depressive symptoms

	Perceived social support			All social s	support sul	oscales	All social support and family criticism subscales		
Step 1 R^2 Step 2 ΔR^2	$R^2 = 0.283$ $\Delta R^2 = 0.008^{**}$			$R^2 = 0.282$ $\Delta R^2 = 0.012^*$				$R^2 = 0.294$ $R^2 = 0.009$	
Step 1	t-value	Un β (SE)		t-value	Un β	(SF)	t-value	Un β	(SE)
Gender	-5.03***	-2.15 (0.43)		-4.95***	-2.17	(0.44)	-5.24***	-2.31	(0.44)
Age	-1.68	-0.05	(0.03)	-1.51	-0.05	(0.03)	-1.09	-0.04	(0.03)
Race	-1.61	-0.43	(0.27)	-1.54	-0.47	(0.30)	-1.56	-0.46	(0.30)
Residence	0.12	0.02	(0.13)	0.11	0.02	(0.13)	0.56	0.08	(0.14)
Education	-1.37	-0.12	(0.09)	-1.00	-0.09	(0.09)	-0.15	-0.01	(0.09)
Cognitive status	0.60	0.06	(0.09)	0.98	0.09	(0.10)	1.43	0.14	(0.10)
CIRS	4.45***	0.36	(0.08)	4.20***	0.35	(0.08)	4.10***	0.35	(0.08)
FEICS	_	_		_			2.42*	0.18	(0.08)
Social interaction	_	_	_	-2.07*	-0.29	(0.14)	-1.95*	-0.28	(0.14)
Perceived social support	-7.81***	-0.75	(0.10)	-6.73***	-0.71	(0.11)	-5.66***	-0.71	(0.13)
Instrumental support		_	<u> </u>	-1.00	-0.09	(0.09)	-0.78	-0.07	(0.09)
Karnofsky	4.93***	0.10	(0.02)	4.56***	0.10	(0.02)	4.88***	0.10	(0.02)
Step 2			()			()			()
Gender	-5.15***	-2.19	(0.43)	-5.06***	-2.21	(0.44)	-5.31***	-2.34	(0.44)
Age	-1.96	-0.06	(0.03)	-1.78	-0.06	(0.03)	-1.28	-0.04	(0.03)
Race	-1.51	-0.40	(0.26)	-1.35	-0.41	(0.30)	-1.28	-0.38	(0.30)
Residence	0.42	0.06	(0.13)	0.35	0.05	(0.13)	0.64	0.09	(0.14)
Education	-1.44	-0.13	(0.09)	-0.99	-0.09	(0.09)	-0.07	-0.01	(0.09)
Cognitive status	0.56	0.05	(0.09)	1.05	0.10	(0.10)	1.25	0.12	(0.10)
CIRS	4.40***	0.36	(0.08)	4.18***	0.35	(0.08)	4.11***	0.35	(0.08)
FEICS	—	—	_	—			2.12*	0.16	(0.08)
Social interaction	—	—	—	-2.01*	-0.28	(0.14)	-1.80	-0.25	(0.14)
Perceived social support	-8.32***	-0.85	(0.10)	-7.48***	-0.83	(0.11)	-6.18***	-0.80	(0.13)
Instrumental support	—	—	—	-0.89	-0.08	(0.09)	-0.99	-0.10	(0.10)
Karnofsky	5.30***	0.11	(0.02)	4.61***	0.10	(0.02)	4.77	0.12	(0.02)
Karnofsky \times FEICS	—	—	—	—	—	—	1.92	0.01	(0.01)
Karnofsky × social interaction	—	—	—	-0.08	0.00	(0.01)	-0.02	0.00	(0.01)
Karnofsky × perceived social support	2.78**	0.02	(0.01)	2.25*	0.02	(0.01)	2.36*	0.02	(0.01)
Karnofsky × instrumental support	—	—	_	-1.12	-0.01	(0.01)	-0.42	-0.00	(0.01)

Cognitive status = MMSE; functional impairment = Karnofsky; illness burden = CIRS; depressive symptoms = Hamilton Depression Rating Scale; social interaction, perceived social support, and instrumental support = subscales of the Duke Social Support Index; family criticism subscale of Family Emotional Involvement and Criticism Scale = FEICS.

***p < 0.001.

depressive symptoms (t = 5.14, p < 0.001; standardized $\beta = 0.1$; Un $\beta = 0.36$ (SE = 0.07)).

In a combined model examining illness burden and all social support subscales simultaneously, greater illness burden was significantly associated with more depressive symptoms (t=4.20, p < 0.001; standardized β =0.18; Un β =0.35 (SE=0.08)), and only social interaction was a significant moderator ($F_{(3, 654)}$ =2.19, p=0.088; t=-2.25, p < 0.05; standardized β =-0.08; Un β =-0.10 (SE=0.05); R^2 = not significant). Those with higher levels of social interaction reported fewer depressive symptoms related to illness burden (Table 3). Main effects existed for social interaction (t=-1.97, p<0.05; standardized β =-0.08; Un β =-0.27 (SE=0.14)) and perceived support (t=-6.81, p<0.001; standardized β =-0.26; Un $\beta = -0.74$ (SE = 0.11)), which were associated with less depressive symptoms.

Similarly, in a combined model examining functional impairment and all social support subscales (Table 2), higher functional impairment was significantly associated with more depressive symptoms (t= 4.61, p < 0.001; standardized β = 0.20) and perceived support was a significant moderator ($F_{(3, 654)}$ = 3.70, p = 0.012; t = 2.25, p < 0.05; standardized β = 0.10). Those with greater perceived support reported fewer depressive symptoms associated with functional impairment. Main effects were found for social interaction (t= -2.01, p < 0.05; standardized β = -0.08) and perceived support (t= -7.48, p < 0.001; standardized β = -0.26), which were related to less depressive symptoms.

^{*}p < 0.05.

^{**}*p* < 0.01.

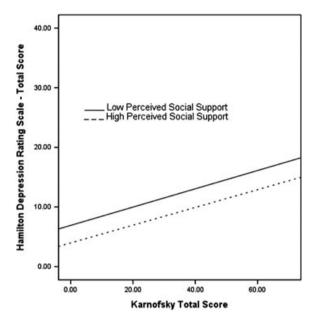


Figure 1 Moderating effect of perceived social support on association between functional impairment and depressive symptoms.

Finally, in a combined model examining illness burden, all social support subscales, and the family criticism scale, illness burden was significantly positively associated with depressive symptoms (t = 3.96, p < 0.001; Un $\beta = 0.34$ (SE = 0.09)), yet only social interaction moderated this relationship (t = -2.43,p < 0.05; standardized $\beta = -0.09$; Un $\beta = -0.11$ (SE = 0.05)) (Table 3). Similarly, in a combined model examining functional impairment and all potential social support and family criticism moderators, only perceived support significantly moderated this relationship (t = 2.36, p < 0.05; standardized $\beta = 0.10$), although there was a trend toward significance for the interaction with family criticism (t=1.92), p = 0.056; standardized $\beta = 0.07$). Individuals with greater perceived support and less perceived family criticism report less depressive symptoms related to functional impairment.

Discussion

Supporting our hypotheses, we found a significant positive association between illness burden, functional impairment and family criticism and depressive symptoms, and a significant negative relationship between perceived social support and social interaction and depressive symptoms. Social interaction significantly buffered the association between illness burden and depressive symptoms, and perceived social support significantly buffered the relationship between functional impairment and depressive symptoms. Neither instrumental support nor family criticism reached significance as moderators. Our results suggest that there may be context-specific differences in the effects of social support; type of support may make a difference, depending on the type of stressor.

For instance, perceived social support was a significant moderator of the association between functional impairment, but not illness burden, and depressive symptoms. Perceived quality of social relations may help older adults cope with deficits in many ways, including increasing life satisfaction, meaningfulness, and feelings of belongingness (Newsom and Schulz, 1996; Park, 2009), and bolstering coping ability (Zautra *et al.*, 2000). When impaired, a sense of the quality and genuineness of social support may be more important for reducing depression risk than simply having many social contacts (Roberson and Lichtenberg, 2003).

On the contrary, social interaction significantly moderated the association between illness burden, but not functional impairment, and depressive symptoms. As an older adult experiences more social interactions, they may feel less lonely and isolated (Street *et al.*, 1999), and more likely to feel supported emotionally. Interactions with others may also facilitate, through encouragement or enhancement of personal control, adaptive health behaviors such as adherence to treatment regimens or engagement in lifestyle changes, which may be beneficial not only for physical but also for mental health (Umberson and Montez, 2010).

Instrumental support was not significantly associated with depressive symptoms, nor was it a moderator, suggesting that older adults experiencing chronic illness and impairment may benefit more, psychologically, from extent and quality of social support (Wallsten *et al.*, 1999). For ill or impaired individuals, receiving instrumental support, while perhaps undesirable, may be viewed as a necessary consequence of one's health status and may not therefore contribute to depression risk. It may be that instrumental support is not as crucial for our sample of ambulatory, older adults presenting for treatment in primary care offices, who are probably not as ill or functionally impaired as those in assisted living and nursing homes (Goodwin and Smyer, 1999).

Family criticism was significantly positively associated with depressive symptoms, confirming previous research (Seaburn *et al.*, 2005), but did not significantly moderate the association between illness burden and depressive symptoms and only neared significance as a buffer of functional impairment. This Table 3 Multivariable hierarchical linear regressions of social relationship variables as moderators of association between illness burden and depressive symptoms

	Perceived social support			All social	support sub	scales	All social support and family criticism subscales			
Step 1 R ²	R ² =0.283				$R^2 = 0.282$			R ² =0.294		
Step 2 ΔR^2		$R^2 = 0.001$			$\Delta R^2 = 0.007$			$R^2 = 0.009$		
Step 1	t-value	Un β		t-value	Un β	· · /	t-value	Un β		
Gender	-4.96***	-2.20	(0.44)	-4.95***	-2.17	(0.44)	-5.24***	-2.31	(0.44)	
Age	-2.04*	-0.07	(0.03)	-1.51	-0.05	(0.03)	-1.09	-0.04	(0.03)	
Race	-2.03*	-0.57	(0.28)	-1.54	-0.47	(0.30)	-1.56	-0.46	(0.30)	
Residence	0.69	0.09	(0.14)	0.11	0.02	(0.13)	0.56	0.08	(0.14)	
Education	-1.54	-0.14	(0.09)	-1.00	-0.09	(0.09)	-0.15	-0.01	(0.09)	
Cognitive status	0.69	0.07	(0.10)	0.98	0.09	(0.10)	1.43	0.14	(0.10)	
Karnofsky	5.85***	0.12	(0.02)	4.56***	0.10	(0.02)	4.88***	0.10	(0.02)	
CIRS	4.33***	0.37	(0.09)	4.20***	0.35	(0.08)	4.10***	0.35	(0.08)	
FEICS			<u> </u>				2.42*	0.18	(0.08)	
Social interaction	-3.59***	-0.49	(0.14)	-2.07*	-0.29	(0.14)	-1.95	-0.28	(0.14)	
Perceived social support	—		—	-6.73***	-0.71	(0.11)	-5.66***	-0.71	(0.13)	
Instrumental support	—		—	-1.00	-0.09	(0.09)	-0.78	-0.07	(0.09)	
Step 2	105***	0.40	(0.4.4)		0.40	(0.44)	E 00***		(0.44)	
Gender	-4.95***	-2.19	(0.44)	-4.94***	-2.16	(0.44)	-5.20***	-2.28	(0.44)	
Age	-1.81	-0.06	(0.03)	-1.35	-0.04	(0.03)	-0.86	-0.03	(0.03)	
Race	-2.12*	-0.60	(0.28)	-1.59	-0.48	(0.30)	-1.58	-0.47	(0.30)	
Residence	0.70	0.10	(0.14)	0.08	0.01	(0.13)	0.56	0.08	(0.14)	
Education	-1.68	-0.16	(0.09)	-1.10	-0.10	(0.09)	-0.33	-0.03	(0.09)	
Cognitive status	0.71	0.07	(0.10)	1.01	0.10	(0.10)	1.53	0.15	(0.10)	
Karnofsky	5.55***	0.12	(0.02)	4.39***	0.09	(0.02)	4.69***	0.10	(0.02)	
CIRS	4.38***	0.37	(0.09)	4.20***	0.35	(0.08)	3.96***	0.34	(0.09)	
FEICS			<u> </u>	4 07*	0.07	(0,4,4)	2.33*	0.18	(0.08)	
Social interaction	-3.42**	-0.47	(0.14)	-1.97*	-0.27	(0.14)	-1.91	-0.27	(0.14)	
Perceived social support	—		—	-6.81***	-0.74	(0.11)	-5.78***	-0.74	(0.13)	
Instrumental support	—		—	-1.02	-0.09	(0.09)	-0.95	-0.09	(0.10)	
			(0, 0, 4)			(0.05)	1.38	0.03	(0.02)	
CIRS × social interaction	-2.50*	-0.11	(0.04)	-2.25*	-0.10	(0.05)	-2.43*	-0.11	(0.05)	
CIRS × perceived social	—	—	—	1.55	0.05	(0.03)	1.23	0.04	(0.04)	
Support CIRS \times instrumental support	—	—	—	-0.48	-0.01	(0.03)	-0.33	-0.01	(0.03)	

Cognitive status = MMSE; functional impairment = Karnofsky; illness burden = CIRS; depressive symptoms = Hamilton Depression Rating Scale; social interaction, perceived social support, and instrumental support = Subscales of the Duke Social Support Index; family criticism subscale of Family Emotional Involvement and Criticism Scale = FEICS.

***p* < 0.01.

***p < 0.001.

lack of findings may be owing to the interpersonal characteristics of our sample, which reported generally high levels of perceived social support and low levels of family criticism. Furthermore, several questions on the FEICS (e.g., *my family finds fault with my friends* and *my family approves of my friends*) may not be suitable for older adults. It may also be the case that already-elevated levels of stress associated with health difficulties supersede distress resulting from familial verbal criticism and, thus, do not additionally compound depression risk.

Our findings may have important implications for interventions in primary care settings that, because of the negative perception of older adults regarding utilization of mental health services (Bogner *et al.*, 2009), are an important "point of capture" (Unützer, 2002). Behavioral health consultants working in primary care should attempt to enhance both the number and quality of social interactions of older patients (Lebowitz et al., 1997). Older adults experiencing chronic illness may benefit from experiential or meaning-based social activities, perhaps involving close family or friends (Jopp and Hertzog, 2010), or participation in religious or spiritual activities (Thune-Boyle et al., 2006). Participation in psychoeducational or self-help groups (Laitinen et al., 2006), or engagement in both leisure and productive social activities, such as physical exercise or attending a senior citizens center (Herzog et al., 1998), may help ill or impaired older adults develop a sense of competency and capability that, in turn, may reduce depressive symptoms.

^{*}p < 0.05.

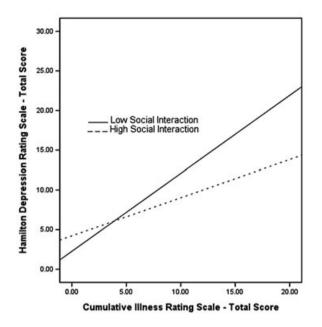


Figure 2 Moderating effect of social interaction on association between illness burden and depressive symptoms.

Older adults with functional impairment, however, may not be willing or able to engage in frequent social interactions, because of decreased mobility and embarrassment or stigma regarding their impairment (Bahm and Forchuk, 2009). In this case, interventions must be focused on either enhancing quality of support, perhaps from caregivers, or assisting the patient in ascribing meaning to the care they receive. Reflective listening techniques, cognitive reframing, or narrative therapy may help a patient to explore themes of loss and devaluation (Kropf and Tandy, 1998). In general, facilitation of interpersonal functioning, or enhancement of the perceived quality of support, may be accomplished using evidence-based treatments, such as cognitive behavioral therapy, which are cost-effective and able to be broadly disseminated (Buenaver et al., 2006).

Despite our study's many strengths, minor limitations must be addressed. Our cross-sectional design prevents the establishment of causal relationships, and bi-directionality of associations is a possibility. Symptoms of depression probably complicate health functioning, and as well, illness and impairment may affect the receipt of social support and family criticism (Fiscella and Campbell, 1999; Meeks *et al.*, 2000); prospective research is needed. A general lack of diversity in our sample limits generalizability, and a more thorough examination of the association between socio-cultural factors and the variables in our study is needed (Musil *et al.*, 1998). In this secondary analysis, the measure of family criticism available, the FEICS, may not be ideally suited for older adults, and our measure of social support did not include an assessment of emotional support, which is of importance to older adult well-being (Oxman *et al.*, 1992). There may also be potentially important variables that we were unable to include in analytic models. Finally, overall, our analytic models only accounted for a small percentage of variance, suggesting that there are other variables that should be accounted for to better explain the relationships posited in our hypotheses. Future research utilizing more appropriate and comprehensive measures, and including variables such as financial strain, loneliness, isolation, helplessness, and hopelessness, is needed.

Despite these limitations, our results suggest that perceived satisfaction with social support and frequency of social interactions are important to the psychological well-being of older adults with chronic illness and functional impairment. Although such health difficulties may have a critical impact on quality of life via the stressful experience of loss of role status, impaired sense of well-being, and disruption of social support (Newsom and Schulz, 1996), our findings suggest that it may be possible to lessen such adverse psychological consequences by increasing perceived social support and promoting social interactions.

Conflict of interest

None declared.

Key points

- Illness burden and functional impairment contribute to risk for depressive symptoms.
- Social support and family criticism affect the associations between health status and depressive symptoms, but their effect may differ depending on nature of the stressor.
- For illness burden, social interaction, and for functional impairment, perceived social support, moderated the association with depressive symptoms.

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