Mathematical Statistics 1

Chapter 2. Multivariate Distributions

2.3. Conditional Distributions and Expectations—Proofs of Theorems



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Theorem 2.3.1

Theorem 2.3.1. Let (X_1, X_2) be a random vector such that the variance of X_2 is finite. Then

(a)
$$E[E[X_2 | X_1]] = E[X_2]$$
, and
(b) $Var([E[X_2 | X_1]) \le Var(X_2)$.

Proof. We give proofs for the continuous case and leave the discrete case as an exercise.

(a) We have

$$E[X_2] = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} x_2 f(x_1, x_2) \, dx_2 \, dx_1$$

=
$$\int_{-\infty}^{\infty} \left(\int_{-\infty}^{\infty} x_2 \frac{f(x_1, x_2)}{f_1(x_1)} \, dx_2 \right) f_1(x_2) \, dx_1$$

=
$$\int_{-\infty}^{\infty} E[X_2 \mid x_1] f_1(x_1) \, dx_1 = E[E[X_2 \mid X_1]]$$

(notice that $E[X_2 | x_1]$ is a function of x_1).

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(notice that $E[X_2 | x_1]$ is a function of x_1).

Theorem 2.3.1 (continued 1)

Proof (continued). (b) Let $\mu_2 = E[X_2]$, then

$$\begin{aligned} \operatorname{Var}(X_2) &= E[(X_2 - \mu_2)^2] = E[(X_2 - E[X_2 \mid X_1] + E[X_2 \mid X_1] - \mu_2)^2] \\ &= E[(X_2 - E[X_2 \mid X_1] + E[X_2 \mid X_1] - \mu_2)(X_2 - E[X_2 \mid X_1] + E[X_2 \mid X_1] - \mu_2)] \\ &= E[(X_2 - E[X_2 \mid X_1])^2 + (E[X_2 \mid X_1] - \mu_2)(X_2 - E[X_2 \mid X_1]) \\ &+ (X_2 - E[X_2 \mid X_1])(E[X_2 \mid X_1] - \mu_2) + (E[X_2 \mid X_1] - \mu_2)^2] \\ &= E[(X_2 - E[X_2 \mid X_1])^2] + 2E[(E[X_2 \mid X_1] - \mu_2)(X_2 - E[X_2 \mid X_1])] \\ &+ E[(E[X_2 \mid X_1] - \mu_2)^2] \end{aligned}$$

where the last equality holds since E is linear by Theorem 2.1.1.

Theorem 2.3.1 (continued 2)

Proof (continued).

$$2E[(E[X_2 \mid X)1] - \mu_2)(X_2 - E[X_2 \mid X_1])]$$

$$=2\int_{-\infty}^{\infty}\int_{-\infty}^{\infty}(X_2-E[X_2\mid x_1])(E[X_2\mid x_1]-\mu_2)f(x_1,x_2)\,dx_2\,dx_1$$

$$=2\int_{-\infty}^{\infty}(E[X_2]-\mu_2)\left(\int_{-\infty}^{\infty}(x_2-E[X_2\mid x_1])\frac{f(x_1,x_2)}{f_1(x_1)}\,dx_2\right)f_1(x_1)\,dx_1.$$

We have

. . .

$$\int_{-\infty}^{\infty} (x_2 - E[X_2 \mid x_1]) \frac{f(x_1, x_2)}{f_1(x_1)} dx_2$$

=
$$\int_{-\infty}^{\infty} x_2 \frac{f(x_1, x_2)}{f_1(x_1)} dx_2 - E[X_2 \mid x_1] \int_{-\infty}^{\infty} \frac{f(x_1, x_2)}{f_1(x_1)} dx_2$$

=
$$E[X_2 \mid x_1] - EpX_2 \mid x_2](1) = 0,$$

Theorem 2.3.1 (continued 3)

Proof (continued). ... so $2E[(X_2 - E[X_2 | X_1])(E[X_2 | X_1] - \mu_2)] = 0$ and

$$Var(X_2) = E[(X_2 - E[X_2 | X_1])^2] + E[(E[X_2 | X_1] - \mu_2)^2]$$

$$\geq E[(E[X_2 | X_1] - \mu_2)^2] \qquad (*)$$

since $E[(X_2 - E[X_2 | X_1])^2] \ge 0$. Now for random variable X, Var $(X) = E[(X - \mu)^2] = E[(X - E[X])^2]$ (see Definition 1.9.2), so random variable $E[X_2 | X_1]$ has mean $E[E[X_2 | X_1]]$ and by part (a), $E[E[X_2 | X_1]] = E[X_2] = \mu_2$ so that

$$Var(E[X_2 | X_1]) = E[(E[X_2 | X_1] - \mu_2)^2]$$

and hence by (*)

 $Var(X_2) \ge E[(E[X_2 | X_1] - \mu_2)^2] = Var(E[X_2 | X_1]),$

as claimed.

Theorem 2.3.1 (continued 3)

Proof (continued). ... so $2E[(X_2 - E[X_2 | X_1])(E[X_2 | X_1] - \mu_2)] = 0$ and

$$Var(X_2) = E[(X_2 - E[X_2 | X_1])^2] + E[(E[X_2 | X_1] - \mu_2)^2]$$

$$\geq E[(E[X_2 | X_1] - \mu_2)^2] \qquad (*)$$

since $E[(X_2 - E[X_2 | X_1])^2] \ge 0$. Now for random variable X, Var $(X) = E[(X - \mu)^2] = E[(X - E[X])^2]$ (see Definition 1.9.2), so random variable $E[X_2 | X_1]$ has mean $E[E[X_2 | X_1]]$ and by part (a), $E[E[X_2 | X_1]] = E[X_2] = \mu_2$ so that

$$Var(E[X_2 | X_1]) = E[(E[X_2 | X_1] - \mu_2)^2]$$

and hence by (*)

$$\mathsf{Var}(X_2) \geq E[(E[X_2 \mid X_1] - \mu_2)^2] = \mathsf{Var}(E[X_2 \mid X_1]),$$

as claimed.