

## Section 1.5. Random Variables

**Note.** We now assign a numerical value to each element  $c$  of the sample space  $\mathcal{C}$ . We use this assignment (called a “random variable”) to associate probabilities with the events (i.e., subsets) in  $\mathcal{C}$ .

**Definition 1.5.1.** Consider a random experiment with a sample space  $\mathcal{C}$ . A function  $X$  which assigns to each  $c \in \mathcal{C}$  one and only one real number  $X(c) = x$  is a *random variable*. The *space* (or *range*) of  $X$  is the set of real numbers  $\mathcal{D} = \{x \mid x = X(c) \text{ for some } c \in \mathcal{C}\}$ . If  $\mathcal{D}$  is a countable set then  $X$  is a *discrete random variable* and if  $\mathcal{D}$  is an interval of real numbers then  $X$  is a *continuous random variable*.

**Note.** We discuss discrete random variables in more detail in the next section (Section 1.6) and discuss continuous random variables in Section 1.7. We illustrate some special cases now.

**Note/Definition.** Suppose  $X$  is a discrete random variable with a finite space (i.e., range)  $\mathcal{D} = \{d_1, d_2, \dots, d_m\}$ . Define the function  $p_X : \mathcal{D} \rightarrow [0, 1]$  as  $p_X(d_i) = P(\{c \mid X(c) = d_i\})$  for  $i = 1, 2, \dots, m$ . In Section 1.6,  $p_X$  is called the *probability mass function* of  $X$ . The induced probability distribution,  $P_X$ , mapping subsets of  $\mathcal{D}$  to  $[0, 1]$  is  $P_X(D) = \sum_{d_i \in D} p_X(d_i)$  for  $D \subset \mathcal{D}$ . Exercise 1.5.11 shows that  $P_X(D)$  actually *is* a probability on  $\mathcal{D}$ .

**Example 1.5.1. Rolling Two Dice.** Let  $X$  be the sum of the surfaces on a roll of a pair of fair 6-sided dice with faces labeled 1 through 6. The sample space is the set of ordered pairs (we assume one die can be distinguished from the other)  $\mathcal{C} = \{(i, j) \mid 1 \leq i, j \leq 6\}$ . Since the dice are fair,  $P(\{(i, j)\}) = 1/36$ . The 36 outcomes are:

$$\begin{array}{cccccc} (1, 1) & (1, 2) & (1, 3) & (1, 4) & (1, 5) & (1, 6) \\ (2, 1) & (2, 2) & (2, 3) & (2, 4) & (2, 5) & (2, 6) \\ (3, 1) & (3, 2) & (3, 3) & (3, 4) & (3, 5) & (3, 6) \\ (4, 1) & (4, 2) & (4, 3) & (4, 4) & (4, 5) & (4, 6) \\ (5, 1) & (5, 2) & (5, 3) & (5, 4) & (5, 5) & (5, 6) \\ (6, 1) & (6, 2) & (6, 3) & (6, 4) & (6, 5) & (6, 6). \end{array}$$

Define random variable  $X$  on the elements of  $\mathcal{C}$  as  $X((i, j)) = i + j$  (so the space is  $\mathcal{D} = \{2, 3, \dots, 11, 12\}$ ). that the probability mass function satisfies:

$x$	2	3	4	5	6	7	8	9	10	11	12
$p_X(x)$	1/36	2/36	3/36	4/36	5/36	6/36	5/36	4/36	3/36	2/36	1/36

Suppose  $B_1 = \{x \mid x = 7, 11\} = \{7, 11\}$  and  $B_2 = \{2, 3, 12\}$  (notice that  $B_1$  and  $B_2$  are subset of  $\mathcal{D}$ ). Then

$$P_X(B_1) = \sum_{x \in B_1} p_X(x) = p_X(7) + p_X(11) = \frac{6}{36} + \frac{2}{36} = \frac{8}{36},$$

$$P_X(B_2) = \sum_{x \in B_2} p_X(x) = p_X(2) + p_X(3) + p_X(12) = \frac{1}{36} + \frac{2}{36} + \frac{1}{36} = \frac{4}{36}.$$

**Note/Definition.** In practice, when  $X$  is a continuous random variable, it often represents measurements (the book refers to a person's weight as an example; see page 38). In this case, we usually have a nonnegative function  $f_X(x)$  such that for intervals  $(a, b) \in \mathcal{D}$ , the induced probability distribution of  $X$ ,  $P_X$ , is defined as

$$P_X((a, b)) = P(\{c \in \mathcal{C} \mid a < X(c) < b\}) = \int_a^b f_X(x) dx.$$

Such an  $f_X$  will be called a *probability density function* (or *pdf*) in  $X$  in Section 1.7. Notice that we require  $\int_{\mathcal{D}} f_X(x) dx = 1$  so that the definition of probability is satisfied by  $P_X$ .

**Note.** So, with  $\mathcal{C}$  as the sample space, we have the random variable  $X : \mathcal{C} \rightarrow \mathbb{R}$ , the space (or range)  $\mathcal{D} = \text{range}(X) \subset \mathbb{R}$ , and the probability distribution  $P_X$  mapping subsets of  $\mathcal{D}$  to  $[0, 1]$  where  $P_X$  is defined by summing on investigating the probability denoted mass function  $f_X$  or  $p_X$  over the subset of  $\mathcal{D}$ .

**Example 1.5.2.** Consider the probability density function

$$f_X(x) = \begin{cases} 1 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise.} \end{cases}$$

Then for  $a, b$  with  $0 \leq a < b \leq 1$  we have

$$P_X((a, b)) = \int_a^b f_X(x) dx = \int_a^b 1 dx = b - a,$$

and, in particular,  $P_X(\mathbb{R}) = P_X((0, 1)) = 1$ . Notice that  $P_X$  is defined on all open intervals of real numbers. Since for Riemann integrals,  $\int_a^a f_X(x) dx = 0$  then  $P_x$  is also defined on all single points (where it is zero;  $P_X(\{x\}) = 0$  for

all  $x \in \mathbb{R}$ ). We can then define  $P_X$  for any closed interval  $[a, b] \subset [0, 1]$  also as  $P_X([a, b]) = \int_a^b f_X(x) dx = b - a$ . We want the subsets of  $\mathbb{R}$  (or subsets of  $[0, 1]$ ) to form a  $\sigma$ -field and, if we restrict ourselves to Riemann integrals, then we have problems. A  $\sigma$ -field which contains all open intervals also contains all Borel sets (see my online notes for Real Analysis 1, MATH 5210, on [1.4. Borel Sets](#)). An example of a Borel set is  $S = \{x \in [0, 1] \mid x \in \mathbb{R} \setminus \mathbb{Q}\}$  (the irrational numbers between 0 and 1). We cannot define a Riemann integral of  $f_X$  over set  $S$ . But the Lebesgue integral of  $f_X$  over  $X$  doesn't exist. In fact, if  $S$  is any Borel set then  $P_X(S) = m(S)$  where  $m(S)$  denotes the Lebesgue measure of set  $S$ . More generally, if  $S$  is any measurable set (the Borel sets are a proper subset of the Lebesgue measurable sets) then  $P_X(S) = m(S)$ ). This discussion is evidence as to why measure theory and Lebesgue integration play such an important role in probability and statistics.

**Definition 1.5.2.** Let  $X$  be a random variable. Its *cumulative distribution function* (or *cdf*) mapping  $F_X : \mathbb{R} \rightarrow [0, 1]$  is

$$F_X(x) = P_X((-\infty, x]) = P(\{x \in \mathcal{C} \mid X(c) \leq x\}) = P(X \leq x).$$

**Exercise 1.5.4.** In Example 1.5.2 we have the probability density function

$$f_X(x) = \begin{cases} 1 & \text{if } 0 < x < 1 \\ 0 & \text{otherwise.} \end{cases}$$

If  $x \leq 0$  then  $P(X \leq x) = \int_{-\infty}^x f_X(t) dt = \int_{-\infty}^x 0 dt = 0$ , if  $0 < x < 1$  then  $P(X \leq x) = \int_{-\infty}^x f_X(t) dt = \int_{-\infty}^0 f_X(t) dt + \int_0^x f_X(t) dt = \int_{-\infty}^0 0 dt + \int_0^x 1 dt = x$ , and

if  $x \geq 1$  then  $P(X \leq x) = \int_{-\infty}^x f_X(t) dt = \int_{-\infty}^0 f_X(t) dt + \int_0^1 f_X(t) dt + \int_1^x f_X(t) dt = \int_{-\infty}^0 0 dt + \int_0^1 1 dt + \int_1^x 0 dt = 0 + 1 + 0 = 1$ . So the cumulative distribution function is

$$F_X(x) = P_X((-\infty, x]) = \int_{-\infty}^x f_X(t) dt = \begin{cases} 0 & \text{if } x < 0 \\ x & \text{if } 0 \leq x < 1 \\ 1 & \text{if } x \geq 1. \end{cases}$$

Notice that  $\frac{d}{dx}[F_X(x)] = \frac{d}{dx} \left[ \int_{-\infty}^x f_X(t) dt \right] = f_X(x)$  (except at  $x = 0$  and  $x = 1$  where  $f_X$  is not differentiable) by the Fundamental Theorem of Calculus.

**Definition.** Let  $X$  and  $Y$  be random variables. Then  $X$  and  $Y$  are *equal in distribution*, denoted  $X \stackrel{D}{=} Y$ , if the cumulative distribution functions are equal,  $F_X(x) = F_Y(x)$  for all  $x \in \mathbb{R}$ .

**Note.** We can have  $X \stackrel{D}{=} Y$  even though  $X$  and  $Y$  are different. With  $Y = 1 - X$  where  $X$  is based on the probability density function of Example 1.5.4 we have  $X \neq Y$  and (1) for  $y < 0$ ,

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(1 - X \leq y) = P(X \geq 1 - y) \\ &= P(X \geq 1 - y) = 1 - P(X < 1 - y) = 1 - (1) \text{ since } 1 - y > 1 \\ &= 0, \end{aligned}$$

(2) for  $y \geq 1$ ,

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(1 - X \leq y) = P(X \geq 1 - y) \\ &= P(X \geq 1 - y) = 1 - P(X < 1 - y) = 1 - (0) \text{ since } 1 - y \leq 0 \end{aligned}$$

$$= 1,$$

and (3) for  $0 \leq y < 1$ ,

$$\begin{aligned} F_Y(y) &= P(Y \leq y) = P(1 - X \leq y) = P(X \geq 1 - y) \\ &= P(X \geq 1 - y) = 1 - P(X < 1 - y) = 1 - (1 - y) \text{ since } 0 < 1 - y \leq 1 \\ &= y. \end{aligned}$$

$$\text{So } F_Y(y) = P(Y \leq y) = \begin{cases} 0 & \text{if } y < 0 \\ y & \text{if } 0 \leq y < 1 \\ 1 & \text{if } y \geq 1. \end{cases} \text{ and hence } X \stackrel{D}{=} Y.$$

**Note.** The following theorem gives some properties of the cumulative density function of a random variable.

**Theorem 1.5.1.** Let  $X$  be a random variable with cumulative distribution function  $F(x)$ . Then

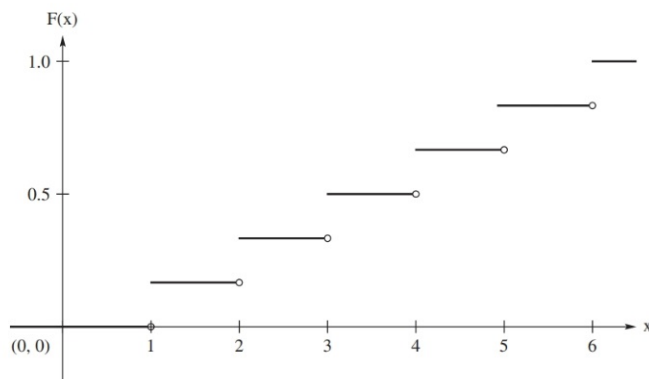
(a) For all  $a$  and  $b$ , if  $a < b$  then  $F(a) \leq F(b)$  (i.e.,  $F$  is nondecreasing).

(b)  $\lim_{x \rightarrow -\infty} F(x) = 0$ .

(c)  $\lim_{x \rightarrow \infty} F(x) = 1$ .

(d)  $\lim_{x \downarrow x_0} F(x) = \lim_{x \rightarrow x_0^+} F(x) = F(x_0)$  (i.e.,  $F$  is right continuous).

**Note.** Theorem 1.5.1(d) does not hold for  $x \uparrow x_0$  or  $x \rightarrow x_0^-$ . That is, the cumulative distribution function need not be continuous from the left. In fact, the cumulative distribution function of the experiment of rolling a fair 6-sided die is not continuous from the right, as shown in Figure 1.5.1.



**Figure 1.5.1.** The cdf for the random variable associated with rolling a fair 6-sided die.

**Note.** The next formula relates certain probabilities to the cumulative distribution function.

**Theorem 1.5.2.** Let  $X$  be a random variable with cumulative distribution function  $F_X$ . Then for  $a < b$  we have  $P(a < X \leq b) = F_X(b) - F_X(a)$ .

**Note.** By Theorem 1.5.1(a), we see that a cumulative distribution function is non-decreasing (i.e., “monotone increasing”) and by definition a cdf function is defined on  $\mathbb{R}$ . Such a function can have at most a countable number of discontinuities (see my online Analysis 1, MATH 4127/5127, notes on [4.2. Monotone and Inverse Functions](#); see Theorem 4-14). The next theorem shows that the discontinuities of the cdf “has mass”; that is, if  $F_X$  is discontinuous at  $x$  then  $P(X = x) > 0$ .

**Theorem 1.5.3.** For random variable  $X$ ,  $P(X = x) = F_X(x) - F_X(x^-)$  for all  $x \in \mathbb{R}$ , where  $F_X(x^-) = \lim_{z \rightarrow x^-} F_X(z)$ .

**Example 1.5.5.** Let  $X$  be the lifetime in years of a mechanical part. Assume the cdf for  $X$  is

$$F_X(x) = \begin{cases} 0 & \text{for } x < 0 \\ 1 - e^{-x} & \text{for } x \geq 0. \end{cases}$$

If we calculate the probability density function by differentiating  $F_X$  then we get

$$\frac{d}{dx}[F_X(x)] = \begin{cases} e^{-x} & \text{for } x > 0 \\ 0 & \text{for } x < 0 \end{cases}$$

since the derivative is modified at  $x = 0$ . The textbook was Theorem 1.5.3 to justify defining the pdf  $f_X$  at  $x = 0$  as  $f_X(0) = 0$ . However, since the cdf and pdf are related as  $F_X(x) = \int_{-\infty}^x f_X(t) dt$  then we can let  $f_X(0)$  be any value because a Riemann integral is unaffected by the value of the integrand at a single point (a Lebesgue integral is also unaffected; however Riemann-Stieltjes and Lebesgue-Stieltjes integrals can be affected by single values of the integrand; see [6.3. The Riemann-Stieltjes Integral](#) from Analysis 2, MATH 4127/5127, and [20.3. Cumulative Distribution Functions and Borel Measures on  \$\mathbb{R}\$](#)  from Real Analysis 2, MATH 5220).

**Note.** We further explore random variables, pdf's, and cdf's in the next two sections.