## Graph Theory

## Chapter 13. The Probabilistic Method

 13.2. Expectation-Proofs of Theorems

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## Lemma 13.1

Lemma 13.1. The Crossing Lemma.
Let $G$ be a simple graph with $m \geq 4 n$. Then $\operatorname{cr}(G) \geq \frac{1}{64} \frac{m^{3}}{n^{2}}$.
Proof. Consider a planar embedding $\tilde{G}$ of $G$ with $\operatorname{cr}(G)$ crossings. Let $S$ be a random subset of $V$ obtained by choosing each vertex of $G$ independently with probability $p=4 n / m \leq 1$. Define the induced subgraph $H=G[S]$ and define the planar embedding $\tilde{H}=\tilde{G}[S]$ (a sub-embedding of $\tilde{G}$ ).

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Define random variables $X, Y, Z$ on $\Omega=\mathcal{P}(V)$ as follows: $X$ is the number of vertices in $S, Y$ is the number of edges in $H=G[S]$, and $Z$ is the number of crossings of $\tilde{H}=\tilde{G}[S]$. By the "trivial lower bound" of Note 13.2.A, for any $S \in \Omega$ we have $Z(S) \geq Y(S)-3 X(S)$.

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## Lemma 13.1 (continued)

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Let $G$ be a simple graph with $m \geq 4 n$. Then $\operatorname{cr}(G) \geq \frac{1}{64} \frac{m^{3}}{n^{2}}$.
Proof (continued). Now $E(X)=p n, E(Y)=p^{2} m$ (since the probability of a particular edge being chosen in the probability that both ends of the edge are chosen), and $E(Z)=p^{4} \operatorname{cr}(G)$ (since each crossing in $\tilde{G}$ is determined by two particular edges or their four particular vertices). Hence

$$
p^{4} \operatorname{cr}(G) \geq p^{2} m-3 p n
$$

or

$$
\operatorname{cr}(G) \geq \frac{m}{p^{2}}-\frac{3 n}{p^{3}}=\frac{p m-3 n}{p^{3}}=\frac{4 n-3 n}{(4 n / m)^{3}}=\frac{n m^{3}}{64 n^{3}}=\frac{1}{64} \frac{m^{3}}{n^{2}},
$$

as claimed.

## Theorem 13.2

Theorem 13.2. Let $P$ be a set of $n$ points in the plane, and let $\ell$ be the number of lines in the plane passing through at least $k+1$ of these points, where $1 \leq k \leq 2 \sqrt{2 n}$. Then $\ell<32 n^{2} / k^{3}$.

Proof. Form a graph $G$ with vertex set $P$ whose edges are the segments between consecutive points on the lines which pass through at least $k+1$ points of set $P$. So each line though $k+1$ points results in a path in $G$ with $k$ edges (and $k+1$ vertices). Since there are $\ell$ such lines, then $G$ has at least $\ell k$ edges.

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## Theorem 13.2 (continued)

Proof (continued). So either $k \ell<4 n$ (in which case The Crossing Lemma does not apply) or $\binom{\ell}{2} \geq \operatorname{cr}(G) \geq \frac{1}{64} \frac{m^{3}}{n^{2}} \geq \frac{1}{64} \frac{(k \ell)^{3}}{n^{2}}$ by The Crossing Lemma. If $k l<4 n$ then

$$
\begin{aligned}
\ell< & \frac{4 n}{k} \leq \frac{4 n}{k}\left(\frac{8 n}{k^{2}}\right) \text { since we hypothesize that } \\
& 1 \leq \frac{2 \sqrt{2 n}}{k} \text { or } 1 \leq \frac{8 n}{k^{2}} \\
= & 32 n / k^{3}, \text { as claimed. }
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If $\binom{\ell}{2} \geq \operatorname{cr}(G) \geq \frac{1}{64} \frac{(k \ell)^{3}}{n^{2}}$ then we have


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If $\binom{\ell}{2} \geq \operatorname{cr}(G) \geq \frac{1}{64} \frac{(k \ell)^{3}}{n^{2}}$ then we have
$\frac{\ell^{2}}{2}>\frac{\ell^{2}-\ell}{2}=\binom{\ell}{2} \geq \operatorname{cr}(G) \geq \frac{1}{64} \frac{(k \ell)^{3}}{n^{2}}$ and so $\ell^{2}>\frac{1}{32} \frac{(k \ell)^{3}}{n^{2}}$ or
$\ell<\frac{32 n^{2}}{k^{3}}$, as claimed.

## Theorem 13.3

Theorem 13.3. Let $P$ be a set of $n$ points in the plane, and let $k$ be the number of pairs of points of $P$ at unit distance. Then $k<5 n^{4 / 3}$.

Proof. Draw a unit circle around each point of $P$. Let $n_{i}$ be the number of these circles passing through exactly $i$ points of $P$. Each circle passes through between 0 and $n-1$ points of $P$ so the total number of circles, $n$, satisfies $n=\sum_{i=0}^{n-1} n_{i}$.

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## Theorem 13.3 (continued 1)

Proof (continued). Now form a graph $H$ with vertex set $P$ whose edges are the arcs of the circles between consecutive points on the circles that pass through at least three points of $P$. Then
$e(H)=\sum_{i=3}^{n-1} i n_{i}=\sum_{i=0}^{n-1} i n_{i}-\left(0 n_{0}+1 n_{1}+2 n_{2}\right)=2 k-n_{1}-2 n_{2} \geq 2 k-2 n$.
Some pairs of vertices of $H$ might be joined by two parallel edges. Delete from $H$ one of each pair of parallel edges, so as to obtain a simple graph $G$ with $e(G) \geq k-n$. Now any two circles in the plane can intersect in at most two points, and since we have $n$ circles then the greatest possible number of intersections of circles is $2 \times\binom{ n}{2}=n(n-1)$. Hence for graph $G, \operatorname{cr}(G) \leq n(n-1)$. If $e(G)<4 n$ then $4 n>e(G) \geq k-n$ and $k \leq 5 n \leq 5 n^{4 / 3}$, as claimed.

## Theorem 13.3 (continued 1)

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e(H)=\sum_{i=3}^{n-1} i n_{i}=\sum_{i=0}^{n-1} i n_{i}-\left(0 n_{0}+1 n_{1}+2 n_{2}\right)=2 k-n_{1}-2 n_{2} \geq 2 k-2 n .
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## Theorem 13.3 (continued 2)

Theorem 13.3. Let $P$ be a set of $n$ points in the plane, and let $k$ be the number of pairs of points of $P$ at unit distance. Then $k<5 n^{4 / 3}$.

Proof (continued). If $e(G) \geq 4 n$ then The Crossing Lemma applies so that

$$
\operatorname{cr}(G) \geq \frac{1}{64} \frac{(e(G))^{3}}{n^{2}} \geq \frac{(k-n)^{3}}{64 n^{2}}
$$

Hence (for $n \geq 1$ )

$$
n^{2}>n^{2}-n=n(n-1) \geq \operatorname{cr}(G) \geq \frac{(k-n)^{3}}{64 n^{2}}
$$

or $n^{2 / 3}>\frac{k-n}{4 n^{2 / 3}}$ or $4 n^{4 / 3}>k-n$ or
$k<4 n^{4 / 3}+n \leq 4 n^{4 / 3}+n^{4 / 3}=5 n^{4 / 3}$, as claimed.

## Proposition 13.4

## Proposition 13.4. Markov's Inequality

Let $X$ be a nonnegative finite random variable on probability space $(\Omega, P)$
and $t>0$. Then $P(X \geq t) \leq \frac{E(X)}{t}$.
Proof. We have by the definition of expectation that

as claimed.

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Proof. We have by the definition of expectation that

$$
\begin{aligned}
& E(X)=\sum_{\omega \in \Omega} X(\omega) P(\omega) \geq \sum_{\omega \in \Omega, X(\omega) \geq t} X(\omega) P(\omega) \\
\geq & \sum_{\omega \in \Omega, X(\omega) \geq t} t P(\omega)=t \sum_{\omega \in \Omega, X(\omega) \geq t} P(\omega)=t P(X \geq t),
\end{aligned}
$$

as claimed.

## Corollary 13.5

Corollary 13.5. Let $X_{n}$ be a nonnegative integer-valued random variable in a probability space $\left(\Omega_{n}, P_{n}\right)$ where $n \in \mathbb{N}$. If $E\left(X_{n}\right) \rightarrow 0$ as $n \rightarrow \infty$, then $P\left(X_{n}=0\right) \rightarrow 1$ as $n \rightarrow \infty$.

Proof. Let $X=X_{n}$ and $t=1$ in Markov's Inequality (Proposition 13.4). Then $P\left(X_{n} \geq 1\right) \leq E\left(X_{n}\right)$ for all $n \in \mathbb{N}$. Since we hypothesize $E\left(X_{n}\right) \rightarrow 0$ as $n \rightarrow \infty$, then it follows (by the Sandwich Theorem, say) that $P\left(X_{n} \geq 1\right) \rightarrow 0$ as $n \rightarrow \infty$. Since $P\left(X_{n}=0\right)=1-P\left(X_{n} \geq 1\right)$ then $P\left(X_{n}=0\right) \rightarrow 1$ as $n \rightarrow \infty$, as claimed.

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## Theorem 13.6

Theorem 13.6. A random graph in $\mathcal{G}_{n, p}$ almost surely has stability number at most $\left\lceil 2 p^{-1} \log n\right\rceil$.

Proof. Let $G \in \mathcal{G}_{n, p}$ and let $S$ be a given set of $k+1$ vertices of $G$ where $k \in \mathbb{N}$. Then $S$ is a stable set of $G$ if none of the $\binom{k+1}{2}$ edges joining two vertices of set $S$ are present in $G$. This happens with probability $(1-p)^{\binom{k+1}{2}}$.

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Let $A_{s}$ denote the event that $S$ is a stable set of $G$, and let $X_{S}$ denote the indicator random variable for this event. By equation (13.5) we have

$$
E\left(X_{S}\right)=P\left(X_{S}=1\right)=P\left(A_{S}\right)=(1-p)\binom{(k+1}{2}
$$

Let $X_{n}$ be the number of stable sets of cardinality $k+1$ in $G \in \mathcal{G}_{n, p}$. Since $X_{S}$ is the indicator random variable then $X_{n}=$


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## Theorem 13.6 (continued 1)

Proof (continued). By linearity of expectation,

$$
\begin{aligned}
& E\left(X_{n}\right)=E\left(\sum_{S \subseteq V,|S|=k+1} X_{S}\right)=\sum_{S \subseteq V,|S|=k+1} E\left(X_{S}\right) \\
& =\sum_{S \subseteq V,|S|=k+1}(1-p)^{\binom{k+1}{2}}=\binom{n}{k+1}(1-p)^{\binom{k+1}{2} .}
\end{aligned}
$$

In Exercise 13.2.1 it is to be shown that

$$
\binom{n}{k+1} \leq \frac{n^{k+1}}{(k+1)!} \text { and } 1-p \leq e^{-p} .
$$

So

$$
\begin{align*}
& \left.E(X)=\binom{n}{k+1}(1-p)^{\binom{k+1}{2} \leq \frac{n^{k+1}}{(k+1)!}\left(e^{-p}\right)^{\binom{k+1}{2}}} \begin{array}{c}
=\frac{n^{k+1} e^{-p(k+1) k / 2}}{(k+1)!}=\frac{\left(n e^{-p k / 2}\right)^{k+1}}{(k+1)!}
\end{array} . .13 .7\right)
\end{align*}
$$

## Theorem 13.6 (continued 1)

Proof (continued). By linearity of expectation,

$$
\begin{aligned}
& E\left(X_{n}\right)=E\left(\sum_{S \subseteq V,|S|=k+1} X_{S}\right)=\sum_{S \subseteq V,|S|=k+1} E\left(X_{S}\right) \\
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\end{aligned}
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$$

So

$$
\begin{gather*}
E(X)=\binom{n}{k+1}(1-p)^{\binom{k+1}{2}} \leq \frac{n^{k+1}}{(k+1)!}\left(e^{-p}\right)^{\binom{k+1}{2}} \\
=\frac{n^{k+1} e^{-p(k+1) k / 2}}{(k+1)!}=\frac{\left(n e^{-p k / 2}\right)^{k+1}}{(k+1)!} \tag{13.7}
\end{gather*}
$$

## Theorem 13.6 (continued 2)

Proof (continued). Suppose $k=\left\lceil 2 p^{-1} \log n\right\rceil$. Then $k \geq 2 p^{-1} \log n$ and so $e^{k} \geq e^{2 p^{-1} \log n}=n^{2 / p}$ or $1 \geq n^{2 / p} e^{-k}$ or $1^{p / 2} \geq\left(n^{2 / p} e^{-k}\right)^{p / 2}$ or $1 \geq n e^{-p k / 2}$. So by equation (13.7), $E\left(X_{n}\right) \leq \frac{1}{(k+1)!}$. As $n \rightarrow \infty$, $k \geq \frac{2}{p} \log n \rightarrow \infty$ and hence $E\left(X_{n}\right) \rightarrow 0$ (this conclusion also holds for any value of $k$ larger than $\left\lceil 2 p^{-1} \log n\right\rceil$ ). Now $X_{n}$ is integer valued and nonnegative, so the hypotheses of Corollary 13.5 are satisfied and hence $P\left(X_{n}=0\right) \rightarrow 1$ as $n \rightarrow \infty$. Since $X_{n}$ is the number of stable sets of cardinality $k+1=\left\lceil 2 p^{-1} \log n\right\rceil+1$ in $G$, we have shown that, almost surely, there are 0 such stable sets. So, equivalently, almost surely the maximum size stable set (that is, the stability number) is of cardinality at most $k=\left\lceil 2 p^{-1} \log n\right\rceil$, as claimed.

## Theorem 13.6 (continued 2)

Proof (continued). Suppose $k=\left\lceil 2 p^{-1} \log n\right\rceil$. Then $k \geq 2 p^{-1} \log n$ and so $e^{k} \geq e^{2 p^{-1} \log n}=n^{2 / p}$ or $1 \geq n^{2 / p} e^{-k}$ or $1^{p / 2} \geq\left(n^{2 / p} e^{-k}\right)^{p / 2}$ or $1 \geq n e^{-p k / 2}$. So by equation (13.7), $E\left(X_{n}\right) \leq \frac{1}{(k+1)!}$. As $n \rightarrow \infty$, $k \geq \frac{2}{p} \log n \rightarrow \infty$ and hence $E\left(X_{n}\right) \rightarrow 0$ (this conclusion also holds for any value of $k$ larger than $\left\lceil 2 p^{-1} \log n\right\rceil$ ). Now $X_{n}$ is integer valued and nonnegative, so the hypotheses of Corollary 13.5 are satisfied and hence $P\left(X_{n}=0\right) \rightarrow 1$ as $n \rightarrow \infty$. Since $X_{n}$ is the number of stable sets of cardinality $k+1=\left\lceil 2 p^{-1} \log n\right\rceil+1$ in $G$, we have shown that, almost surely, there are 0 such stable sets. So, equivalently, almost surely the maximum size stable set (that is, the stability number) is of cardinality at most $k=\left\lceil 2 p^{-1} \log n\right\rceil$, as claimed.

