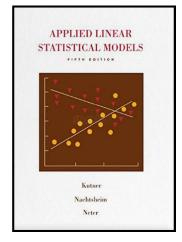
Applied Linear Statistical Models, Part 1

Section 5.8. Random Vectors and Matrices—Proofs of Theorems



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Theorem 5.8.A (continued 1)

Theorem 5.8.A. For Y a random vector and A a matrix of scalars (i.e., constants), we can define the new random variable $\mathbf{W} = \mathbf{AY}$. Then:

(5.45) Expectation of a Constant Matrix Times a Random Vector:

$$E\{W\} = E\{AY\} = AE\{Y\}.$$

Proof (continued). (5.45) Let $A = [a_{ij}]$ be an $m \times n$ matrix of scalars and let $\mathbf{Y} = [Y_i]$ be an $n \times 1$ vector of random variables. Then the *i*th entry of **AY** is $\sum_{i=1}^{n} a_{ij} Y_{j}$. So the *i*th entry of **E**{**AY**} is $\mathbf{E}\left\{\sum_{i=1}^{n} a_{ij} Y_{j}\right\} = \sum_{i=1}^{n} a_{ij} \mathbf{E}\left\{Y_{j}\right\}$, because by Theorem 1.8.2 in my online notes for Mathematical Statistics 1 [STAT 4047/5047] on Section 1.8. Expectation of a Random Variable, expectation is linear. Similarly, the ith entry of $AE\{Y\}$ is $\sum_{i=1}^{n} a_{ij} E\{Y_j\}$ and so $E\{AY\} = AE\{Y\}$, as claimed.

Theorem 5.8.A

Theorem 5.8.A. For **Y** a random vector and **A** a matrix of scalars (i.e., constants), we can define the new random variable $\mathbf{W} = \mathbf{AY}$. Then:

- (5.44) Expectation of a Constant Matrix: $\mathbf{E}\{\mathbf{A}\} = \mathbf{A}$.
- (5.45) Expectation of a Constant Matrix Times a Random Vector:

$$\mathbf{E}\{\mathbf{W}\} = \mathbf{E}\{\mathbf{AY}\} = \mathbf{AE}\{\mathbf{Y}\}.$$

(5.46) Variance of a Constant Matrix Times a Random Vector:

$$\sigma^2\{\mathbf{W}\} = \sigma^2\{\mathbf{AY}\} = \mathbf{A}\sigma^2\{\mathbf{Y}\}\mathbf{A}'.$$

Proof. (5.44) This follows from the definition of the expectation of a matrix and the fact that the expectation of a constant is the constant itself, as claimed.

Theorem 5.8.A (continued 2)

Theorem 5.8.A. For Y a random vector and A a matrix of scalars (i.e., constants), we can define the new random variable $\mathbf{W} = \mathbf{AY}$. Then:

(5.46) Variance of a Constant Matrix Times a Random Vector:

$$oldsymbol{\sigma}^2\{\mathbf{W}\}=oldsymbol{\sigma}^2\{\mathbf{AY}\}=\mathbf{A}oldsymbol{\sigma}^2\{\mathbf{Y}\}\mathbf{A}'.$$

Proof (continued). (5.46) With notation introduced for (5.45), the (i,j)-entry of $\sigma^2\{\mathbf{AY}\}$ is $\mathbf{E}\left\{\left(\sum_{j=1}^n a_{ij} E\{Y_j\} - E\left(\sum_{j=1}^n a_{ij} E\{Y_j\}\right)\right)^2\right\}$. By Note 5.8.A.

$$\sigma^{2}\{\mathbf{W}\} = \mathbf{E}\left\{ [W_{j} - E\{W_{j}\}][W_{j} - E\{W_{j}\}]' \right\}$$

$$= \mathbf{E} \left\{ \left[\sum_{j=1}^{n} a_{ij} Y_j - E \left\{ \sum_{j=1}^{n} a_{ij} Y_j \right\} \right] \left[\sum_{j=1}^{n} a_{ij} Y_j - E \left\{ \sum_{j=1}^{n} a_{ij} Y_j \right\} \right]' \right\}$$

Theorem 5.8.A (continued 3)

Proof (continued). ...

$$\sigma^{2}\{\mathbf{W}\} = \mathbf{E} \left\{ \left[\sum_{j=1}^{n} a_{ij} Y_{j} - \sum_{j=1}^{n} a_{ij} E\{Y_{j}\} \right] \left[\sum_{j=1}^{n} a_{ij} Y_{j} - \sum_{j=1}^{n} a_{ij} E\{Y_{j}\} \right]' \right\}$$

$$= \mathbf{E} \{ (\mathbf{A}[Y_{j} - E\{Y_{j}\}]) (\mathbf{A}[Y_{j} - E\{Y_{j}\}]') \}$$

$$= \mathbf{E} \{ \mathbf{A}[Y_{j} - E\{Y_{j}\}] [Y_{j} - E\{Y_{j}\}]' \mathbf{A}' \} \text{ by (5.32)}$$

$$= \mathbf{A} \mathbf{E} \{ [Y_{j} - E\{Y_{j}\}] [Y_{j} - E\{Y_{j}\}]' \} \mathbf{A}' \text{ since } \mathbf{E} \text{ is linear}$$

$$= \mathbf{A} \sigma^{2} \{ \mathbf{Y} \} \mathbf{A}' \text{ by Note 5.8.A,}$$

as claimed.

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