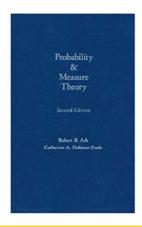
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Chapter 4. Basic Concepts of Probability

4.5. Conditional Probability—Proofs of Theorems



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Theorem 4.5.1 (continued)

Proof (continued). Inductively, we have

$$P((A_1 \cap A_2 \cap \dots \cap A_{n-1}) \cap A_n)$$
= $P((A_1 \cap A_2 \cap \dots \cap A_{n-2}) \cap A_{n-1}) P(A_n \mid A_1 \cap A_2 \cap \dots \cap A_{n-1})$
= $P((A_1 \cap A_2 \cap \dots \cap A_{n-2}) P(A_{n-1} \mid A_1 \cap A_2 \cap \dots \cap A_{n-2})$

$$= P((A_1 \cap A_2 \cap \cdots \cap A_{n-2})P(A_{n-1} \mid A_1 \cap A_2 \cap \cdots \cap A_{n-2}) \times P(A_n \mid A_1 \cap A_2 \cap \cdots \cap A_{n-1})$$

$$= P((A_1 \cap A_2 \cap \cdots \cap A_i)P(A_{i+1} \mid A_1 \cap A_2 \cap \cdots \cap A_i) \times P(A_{i+2} \mid A_1 \cap A_2 \cap \cdots \cap A_{i+1}) \cdots P(A_n \mid A_1 \cap A_2 \cap \cdots \cap A_{n-1})$$

$$= P(A_1)P(A_2 \mid A_1)P(A_3 \mid A_1 \cap A_2) \cdots P(A_{n-1} \mid A_1 \cap A_2 \cap \cdots \cap A_{n-2}) \times P(A_n \mid A_1 \cap A_2 \cap \cdots \cap A_{n-1})$$

Theorem 4.5.1

Theorem 4.5.1.

- (a) Let P(A) > 0. Events A and B are independent if and only if $P(B \mid A) = P(B)$.
- (b) Let $P(A_1 \cap A_2 \cap \cdots \cap A_{n-1}) > 0$. Then

$$P(A_1 \cap A_2 \cap \cdots \cap A_n) = P(A_1)P(A_2 \mid A_1)P(A_3 \mid A_1 \cap A_2) \cdots \cdots P(A_n \mid A_1 \cap A_2 \cap \cdots \cap A_{n-1}).$$

Proof. (a) By definition, $P(A \cap B) = P(B \mid A)P(A)$. If A and B are independent then $P(A \cap B) = P(A)P(B)$ so that we have $P(B) = P(B \mid A)$, as claimed. If $P(B \mid A) = P(B)$, then we have $P(A \cap B) = P(A)P(B)$ and so A and B are independent, as claimed.

(b) Since $P(A_1 \cap A_2 \cap \cdots \cap A_{n-1}) > 0$ then by monotonicity of measure, $P(A_1 \cap A_2 \cap \cdots \cap A_k) > P(A_1 \cap A_2 \cap \cdots \cap A_{n-1}) > 0$ for $k = 1, 2, \ldots, n - 1.$

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Theorem 4.5.2

Theorem 4.5.2. Theorem of Total Probability.

Let B_1, B_2, \ldots form a finite or countably infinite family of mutually exclusive and exhaustive events; that is, $\bigcup_i B_i = \Omega$.

- (a) If A if any event, then $P(A) = \sum_i P(A \cap B_i)$. Thus P(A) is calculated by making a list of mutually exclusive exhaustive ways in which A can happen, and adding the individual probabilities.
- (b) $P(A) = \sum_{i} P(B_i) P(A \mid B_i)$ where the sum is taken over those i for which $P(B_i) > 0$. Thus P(A) is a weighted average of the conditional probabilities $P(A \mid B_i)$.

Proof. (a) We have by countable additivity:

$$P(A) = P(A \cap \Omega) = P(A \cap (\cup_i B_i)) = P(\cup_i (A \cap B_i)) = \sum_i P(A \cap B_i).$$

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Theorem 4.5.2 (continued)

Proof (continued). (b) By part (a), $P(A) = \sum_i P(A \cap B)i$). If $P(B_i) = 0$ for some i then by monotonicity $P(A \cap B_i) \leq P(B_i) = 0$. If $P(B_i) > 0$ then by definition of conditional probability, $P(A \cap B_i) = P(B_i)P(A \mid B_i)$. So

$$P(A) = \sum_{i} P(A \cap B_i) = \sum_{i'} P(B_{i'}P(A \mid B_{i'})$$

where i' ranges over all values for which $P(B_{i'}) > 0$, as claimed.

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