## **Real Analysis**

### **Chapter 4. Basic Concepts of Probability** 4.5. Conditional Probability—Proofs of Theorems



**Real Analysis** 





#### Theorem 4.5.1.

(a) Let P(A) > 0. Events A and B are independent if and only if P(B | A) = P(B).
(b) Let P(A<sub>1</sub> ∩ A<sub>2</sub> ∩ · · · ∩ A<sub>n-1</sub>) > 0. Then P(A<sub>1</sub> ∩ A<sub>2</sub> ∩ · · · ∩ A<sub>n</sub>) = P(A<sub>1</sub>)P(A<sub>2</sub> | A<sub>1</sub>)P(A<sub>3</sub> | A<sub>1</sub> ∩ A<sub>2</sub>) · · · · · · P(A<sub>n</sub> | A<sub>1</sub> ∩ A<sub>2</sub> ∩ · · · ∩ A<sub>n-1</sub>).

**Proof.** (a) By definition,  $P(A \cap B) = P(B \mid A)P(A)$ . If A and B are independent then  $P(A \cap B) = P(A)P(B)$  so that we have  $P(B) = P(B \mid A)$ , as claimed. If  $P(B \mid A) = P(B)$ , then we have  $P(A \cap B) = P(A)P(B)$  and so A and B are independent, as claimed.

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(b) Since  $P(A_1 \cap A_2 \cap \cdots \cap A_{n-1}) > 0$  then by monotonicity of measure,  $P(A_1 \cap A_2 \cap \cdots \cap A_k) > P(A_1 \cap A_2 \cap \cdots \cap A_{n-1}) > 0$  for k = 1, 2, ..., n-1.

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## Theorem 4.5.1 (continued)

Proof (continued). Inductively, we have

$$P((A_1 \cap A_2 \cap \dots \cap A_{n-1}) \cap A_n)$$

$$= P((A_1 \cap A_2 \cap \dots \cap A_{n-2}) \cap A_{n-1})P(A_n \mid A_1 \cap A_2 \cap \dots \cap A_{n-1}))$$

$$= P((A_1 \cap A_2 \cap \dots \cap A_{n-2})P(A_{n-1} \mid A_1 \cap A_2 \cap \dots \cap A_{n-2}))$$

$$\times P(A_n \mid A_1 \cap A_2 \cap \dots \cap A_{n-1})$$

$$= P((A_1 \cap A_2 \cap \cdots \cap A_i)P(A_{i+1} \mid A_1 \cap A_2 \cap \cdots \cap A_i))$$
  
 
$$\times P(A_{i+2} \mid A_1 \cap A_2 \cap \cdots \cap A_{i+1}) \cdots P(A_n \mid A_1 \cap A_2 \cap \cdots \cap A_{n-1})$$

 $= P(A_1)P(A_2 | A_1)P(A_3 | A_1 \cap A_2) \cdots P(A_{n-1} | A_1 \cap A_2 \cap \cdots \cap A_{n-2}) \\ \times P(A_n | A_1 \cap A_2 \cap \cdots \cap A_{n-1})$ 

.

### Theorem 4.5.2. Theorem of Total Probability.

Let  $B_1, B_2, \ldots$  form a finite or countably infinite family of mutually exclusive and exhaustive events; that is,  $\bigcup_i B_i = \Omega$ .

- (a) If A if any event, then  $P(A) = \sum_{i} P(A \cap B_i)$ . Thus P(A) is calculated by making a list of mutually exclusive exhaustive ways in which A can happen, and adding the individual probabilities.
- (b)  $P(A) = \sum_{i} P(B_i)P(A | B_i)$  where the sum is taken over those *i* for which  $P(B_i) > 0$ . Thus P(A) is a weighted average of the conditional probabilities  $P(A | B_i)$ .

**Proof.** (a) We have by countable additivity:

$$P(A) = P(A \cap \Omega) = P(A \cap (\bigcup_i B_i)) = P(\bigcup_i (A \cap B_i)) = \sum_i P(A \cap B_i).$$

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# Theorem 4.5.2 (continued)

**Proof (continued). (b)** By part (a),  $P(A) = \sum_{i} P(A \cap B)i$ ). If  $P(B_i) = 0$  for some *i* then by monotonicity  $P(A \cap B_i) \le P(B_i) = 0$ . If  $P(B_i) > 0$  then by definition of conditional probability,  $P(A \cap B_i) = P(B_i)P(A \mid B_i)$ . So

$$P(A) = \sum_{i} P(A \cap B_i) = \sum_{i'} P(B_{i'}P(A \mid B_{i'}))$$

where i' ranges over all values for which  $P(B_{i'}) > 0$ , as claimed.

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